



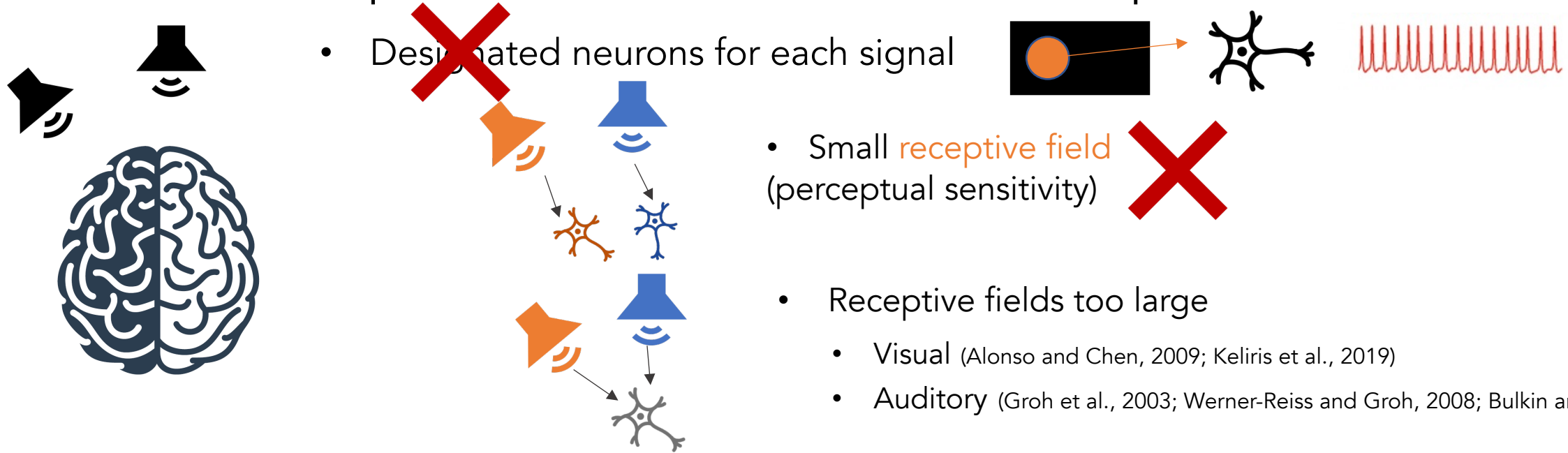
# Detection of Code Juggling with Spike Counts Data Analysis

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# Motivation

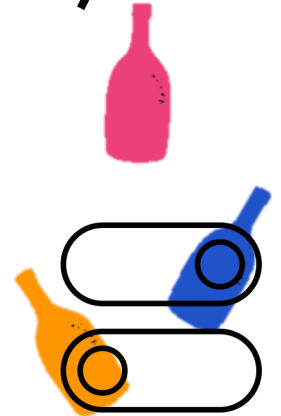
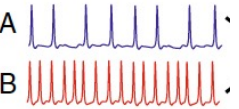
How does brain preserve information about multiple simultaneous items ?



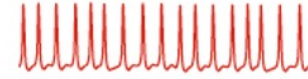
A single neuron will be exposed to multiple simultaneous stimuli !

- Dynamics in presentation in a neural level
- Can a single neuron preserve info from both  and 

# Motivation: potential dynamics



- Always encode A (or B)



- 1<sup>st</sup> order stochasticity
- Poisson distribution  
(Ventura et al., 2002; Kass et al., 2005)

- Treat it as a new stimulus



- Switch between A and B

- **Code Juggling**

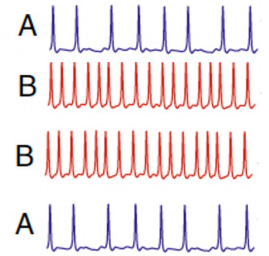
- **2<sup>nd</sup> order stochasticity: constituent stimuli**

- Greater efficiency on information preserving
- Overdispersed + bounded by A & B

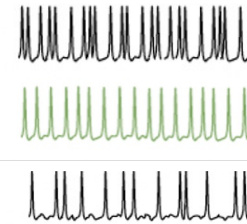
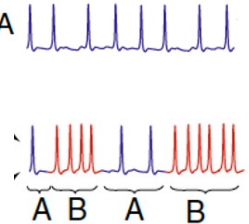
- Stochasticity not related A nor B

- Overdispersed
- Not related to A nor B

- Across trials  
(Caruso et al., 2018)



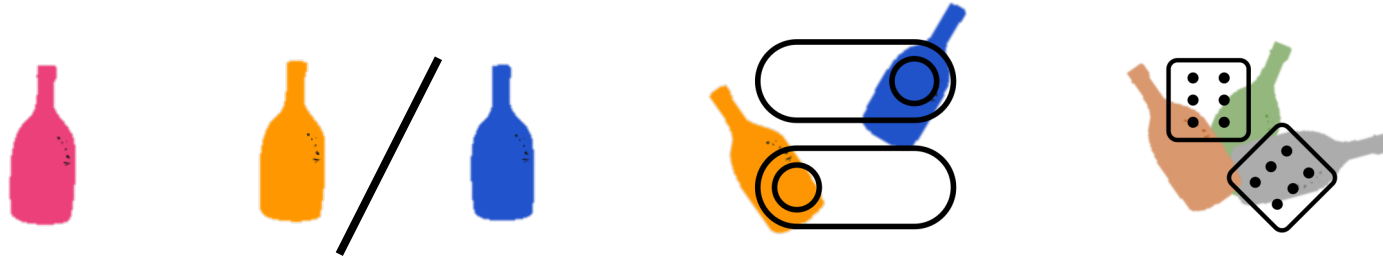
- Within a trial  
(Glynn et al., 2021)



- Higher trial-to-trial variability  
(Semedo et al., 2019)

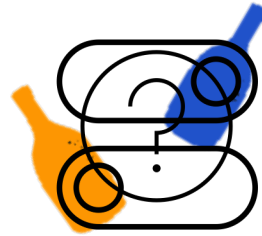
# Research question:

How to **model** these biologically meaningful encoding schemes?



Poisson Mixtures

How to **select** the most likely scheme and quantify the uncertainty?

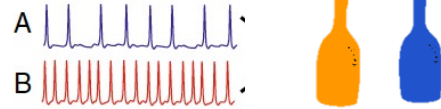


Hypothesis Testing  
(Bayes Factor)

# Statistical Analysis Framework

Assumptions:

- Two-stimuli cases



- Single stimulus spike counts follow Poisson distribution

$$P^A = \text{Poi}(\mu^A), \quad P^B = \text{Poi}(\mu^B), \quad \mu^A, \mu^B > 0.$$



Data:  $(Y_j^e : j \in \{1, \dots, n_e\}, e \in \{A, B, AB\})$

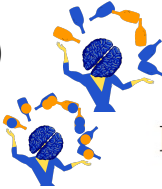
Hypotheses:



FIXED:  $P^{AB} = \text{Poi}(\mu),$

$\mu \in (0, \infty),$

(whole-trial)



MIXTURE:  $P^{AB} = \rho \text{Poi}(\mu^A) + (1 - \rho) \text{Poi}(\mu^B),$

$\rho \in (0, 1),$

(sub-trial)

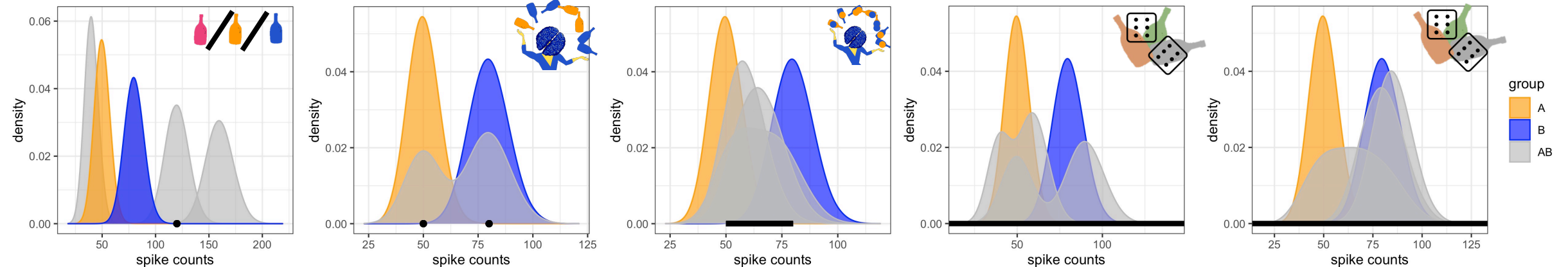
INTERMEDIATE:  $P^{AB} = \int \text{Poi}(r\mu^A + (1 - r)\mu^B) f(r) dr,$   $f$  is a density on  $(0, 1),$



OVERREACHING:  $P^{AB} = \int \text{Poi}(m) g(m) dm,$

$g$  is a density on  $(0, \infty),$

Probability distributions (support + continuity):



# Hypothesis Testing in a Bayesian framework

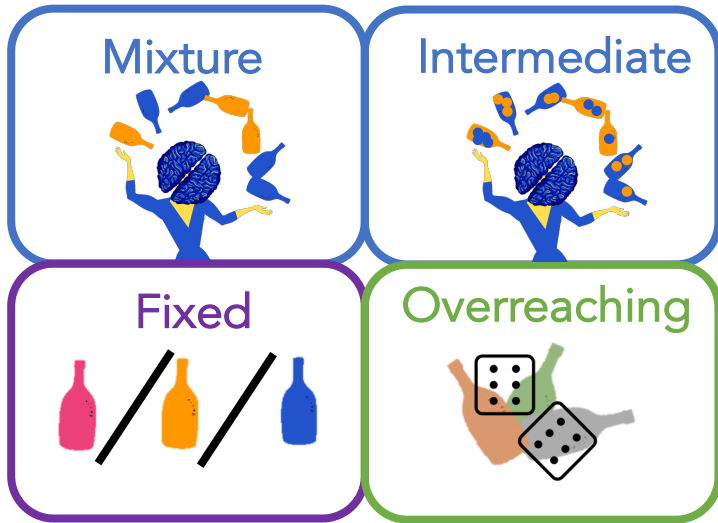
Prior



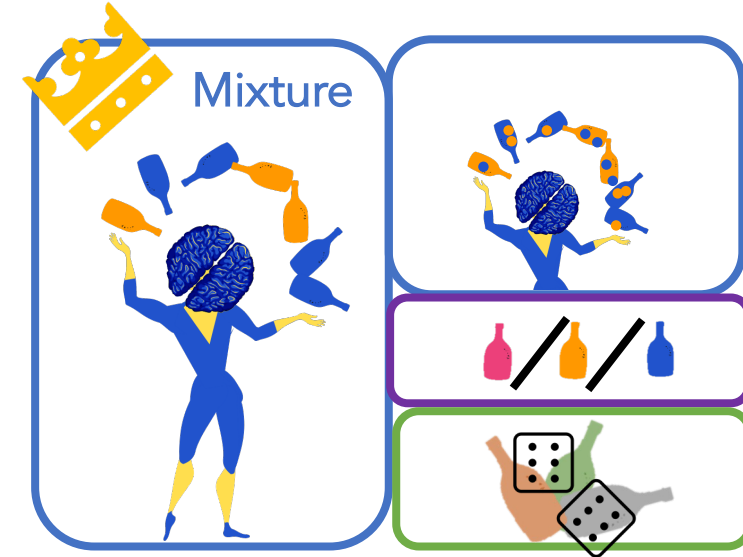
Data



Posterior



Spike counts  
( $Y_j^e : j \in \{1, \dots, n_e\}, e \in \{A, B, AB\}$ )



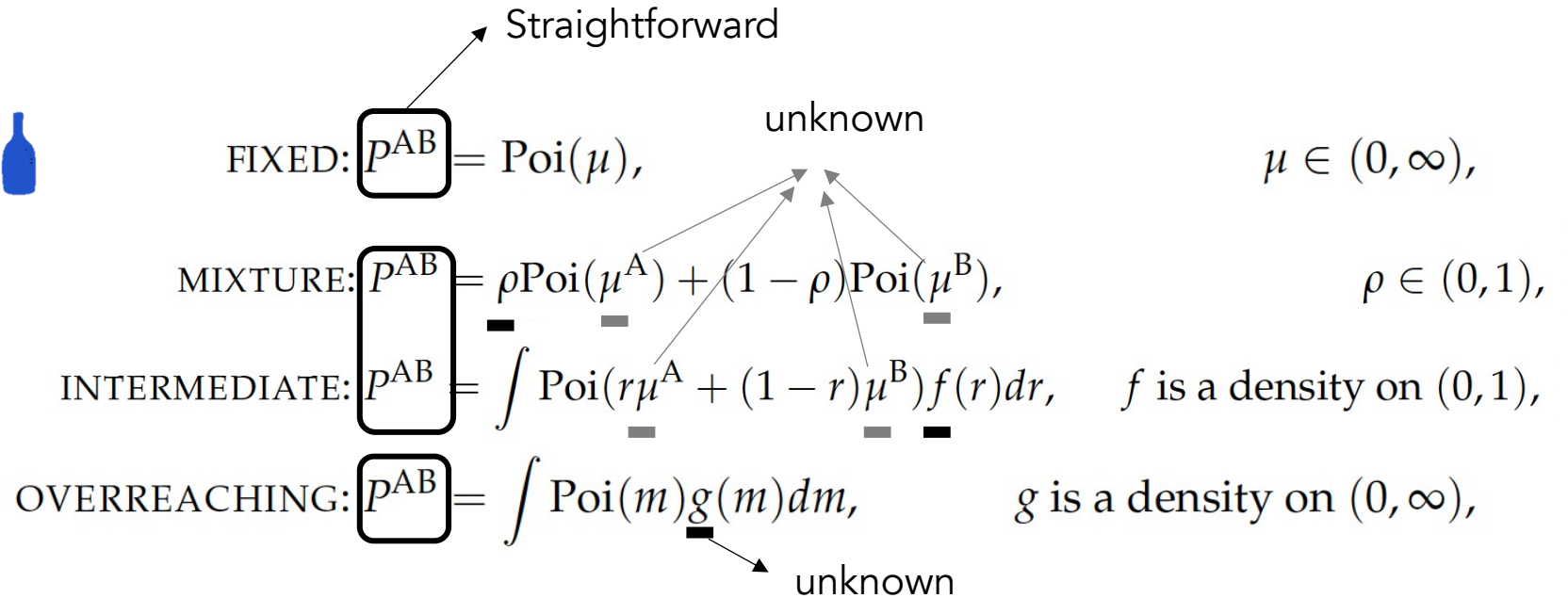
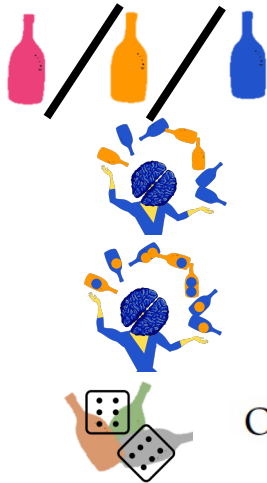
Challenges: nonparametric density estimation and fair competition

Predictive recursion marginal likelihood method (PRML)

(Newton et al., 1998; Newton, 2002; Tokdar et al., 2009; Martin and Tokdar, 2011)

# Statistical Analysis Framework

Hypotheses:



# Statistical Analysis Framework

Hypotheses:



FIXED:  $P^{AB} = \text{Poi}(\mu)$ ,

$\mu \in (0, \infty)$ ,

Marginal likelihood:

$$p(Y^A, Y^B, Y^{AB}) = \cancel{p(Y^A)} \cancel{p(Y^B)} p(Y^{AB} | Y^A, Y^B)$$

$$p(Y^{AB} | Y^A, Y^B) = p(Y^{AB}) = \int p(Y^{AB} | \mu) \pi(\mu) d\mu$$

Set a prior and can obtain closed form



# Statistical Analysis Framework



OVERREACHING:  $P^{AB} = \int \text{Poi}(m)g(m)dm$ ,  $g$  is a density on  $(0, \infty)$ ,

Marginal likelihood:  $p(Y^{AB}|Y^A, Y^B) = p(Y^{AB})$

## Predictive Recursion (Smith and Markov, 1978; Newton et al., 1998; Newton, 2002)

Data  $Y_1, \dots, Y_n$  from mixture distribution

$$m_f(y) = \int \overset{\text{known}}{\kappa(y | u)} \overset{\text{unknown}}{f(u)} d\nu(u)$$

For  $i \in \{1, \dots, n\}$

$$f_i(u) = (1 - w_i)f_{i-1}(u) + w_i \frac{\kappa(Y_i | u)f_{i-1}(u)}{m_{i-1}(Y_i)}, \quad u \in \mathcal{U},$$

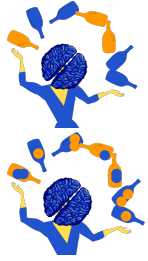
$$m_{i-1}(y) = \int \kappa(y | u)f_{i-1}(u)d\nu(u), \quad y \in \mathcal{Y}$$

Theoretical guarantee (Tokdar et al., 2009) :

$f_n$   $m_n$  converge to the truth asymptotically ( provided  $\sum_{i=1}^{\infty} w_i = \infty$   $\sum_{i=1}^{\infty} w_i^2 < \infty$  )

Bayesian inferential paradigm interpretation (Tokdar et al., 2009)

# Statistical Analysis Framework



$$\text{MIXTURE: } P^{AB} = \rho \text{Poi}(\mu^A) + (1 - \rho) \text{Poi}(\mu^B), \quad \rho \in (0, 1),$$

$$\text{INTERMEDIATE: } P^{AB} = \int \text{Poi}(r\mu^A + (1 - r)\mu^B) f(r) dr, \quad f \text{ is a density on } (0, 1),$$

$$\text{Marginal likelihood: } p(Y^{AB} | Y^A, Y^B) = \int p(Y^{AB} | \theta) p(\theta | Y^A, Y^B) d\theta \quad \theta = (\mu^A, \mu^B)$$

$$\text{First-stage prior: } \pi_0(\mu^A, \mu^B) = 1 / \sqrt{\mu^A \mu^B},$$

$$\text{Second-stage prior: } \pi(\mu^A, \mu^B) = \text{Gam}(\mu^A | 0.5 + \sum_{j=1}^{n_A} Y_j^A, n_A) \times \text{Gam}(\mu^B | 0.5 + \sum_{j=1}^{n_B} Y_j^B, n_B).$$

$$m_f(y) = \int \kappa_\theta(y | u) f(u) dv(u)$$

# Predictive Recursion Marginal Likelihood

Data  $Y_1, \dots, Y_n$  from mixture distribution

$$i \in \{1, \dots, n\}$$

$$m_f(y) = \int \kappa_\theta(y | u) f(u) d\nu(u)$$

$$f_i(u) = (1 - w_i) f_{i-1}(u) + w_i \frac{\kappa_\theta(Y_i | u) f_{i-1}(u)}{m_{i-1}(Y_i)}, \quad u \in \mathcal{U},$$

$$m_{i-1,\theta}(y) = \int \kappa_\theta(y | u) f_{i-1}(u) d\nu(u), \quad y \in \mathcal{Y}$$

$$\kappa = \kappa_\theta$$

$$L_n(\theta) := \prod_{i=1}^n m_{i-1,\theta}(Y_i)$$

## Predictive Recursion

(Newton et al., 1998; Newton, 2002)

## PRML score

(Martin and Tokdar, 2011)

- PRML score estimate true marginal density:  $p(Y_1, \dots, Y_n | \theta) = \prod_{i=1}^n p(Y_i | Y_1, \dots, Y_{i-1}, \theta)$
- Asymptotic consistency:  $n^{-1} \log \{L_n(\theta) / L_n(\theta^*)\} \rightarrow -\inf_{f \in \mathbb{F}} d_{\text{KL}}(m^*, m_{f,\theta})$ .
- Interpretable: expectation filtration approximation to a fully Bayesian estimation  $w_i = (1 + a)^{-1}$ ,  $a > 0$ .

Laplace Approximation: 
$$I(M_h) := \int_{\Theta_h} L_{n,h}(\theta_h) \pi_h(\theta_h) d\theta_h, \quad h \in H,$$

$$\hat{I}(M_h) = L_{n,h}(\hat{\theta}_h) (2\pi)^{-d_h/2} |\Sigma_h|^{1/2},$$

# Fair competition: choice of prior

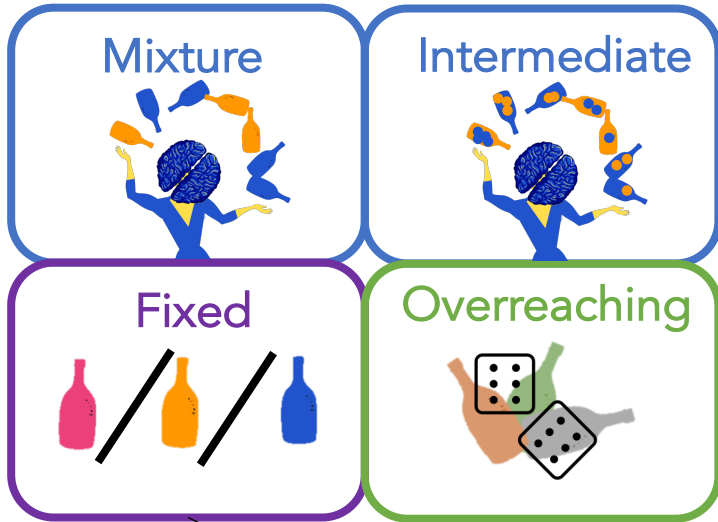
Prior



Data



Posterior



Spike counts

$(Y_j^e : j \in \{1, \dots, n_e\}, e \in \{A, B, AB\})$



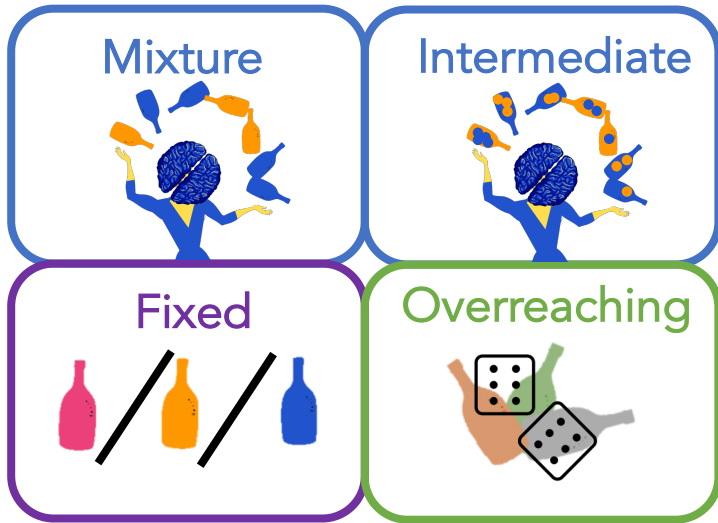
$$p_h(Y) = \frac{I(M_h) p_{0,h}}{\sum_{h' \in H} I(M_{h'}) p_{0,h'}}$$

- Sensitive to the choice of prior
- Bounded uniform (same as overreaching)
- Jeffreys' prior (improper) - - > tune the constant
- Jeffreys' prior (intrinsic Bayes factor) (Berger & Pericchi, 1996)



# Two potential frameworks

## Full Bayesian



Jeffreys' prior: tune constant to match the confidence level

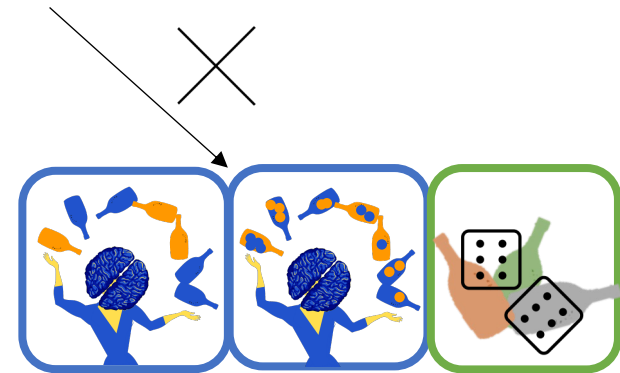
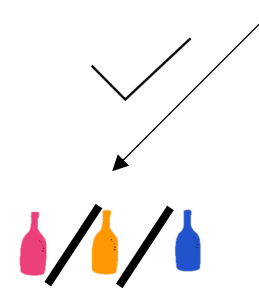
$$p(Y^{AB}|Y^A, Y^B) = p(Y^{AB}) = \int p(Y^{AB}|\mu)\pi(\mu)d\mu$$

$$\pi(\mu) \propto \mu^{-1/2}$$

## Two-stage procedure

Poisson variance test  
(Brown and Zhao, 2002)

Is it Fixed?



# Performance Assessment

100 experiment sets:  $P^A = \text{Poi}(50)$        $P^B = \text{Poi}(80)$        $n_A = n_B = n_{AB}$

FIXED:  $P^{AB} = \text{Poi}(80m)$

MIXTURE:  $P^{AB} = m\text{Poi}(50) + (1 - m)\text{Poi}(80)$

INTERMEDIATE:  $P^{AB} = \int \text{Poi}(r50 + (1 - r)80)\text{Beta}(r; a, b)dr$

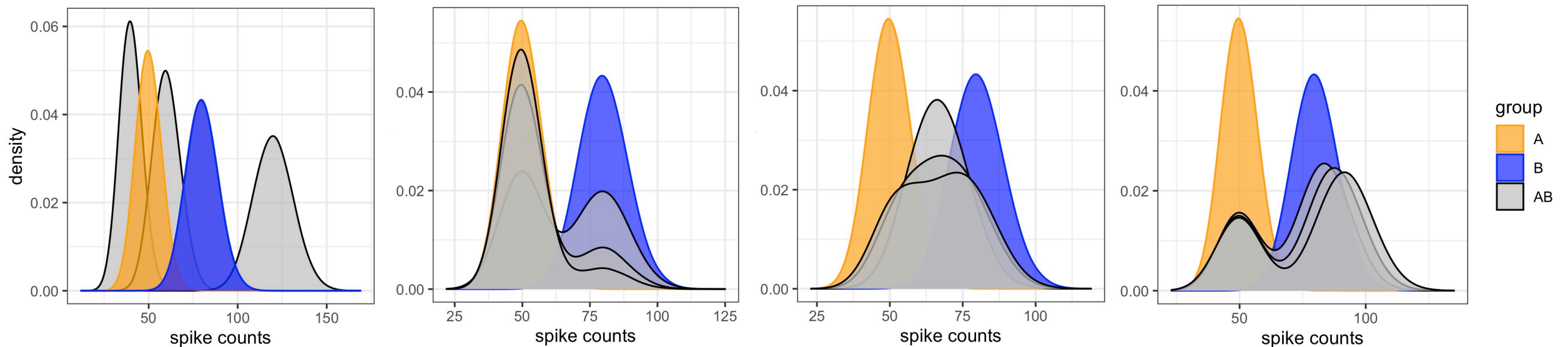
OVERREACHING:  $P^{AB} = 1/3\text{Poi}(50) + 2/3\text{Poi}(80m)$

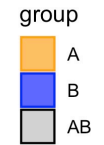
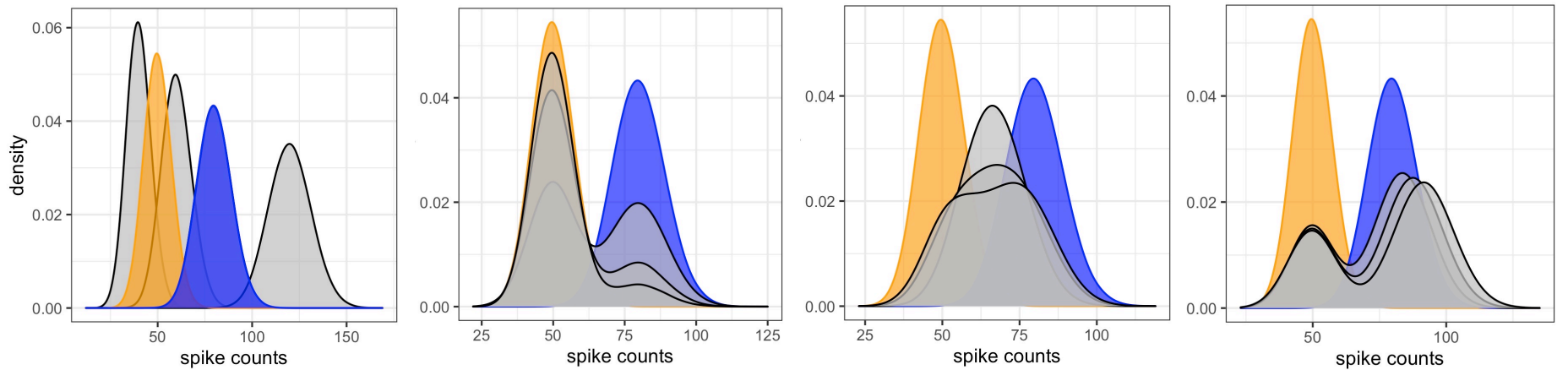
Multiplier

Skewness

Fano factor

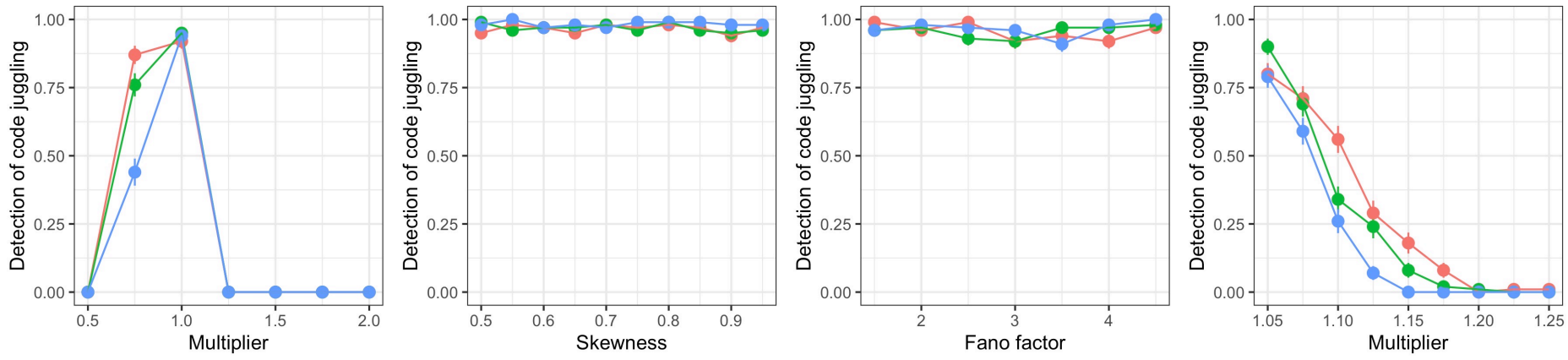
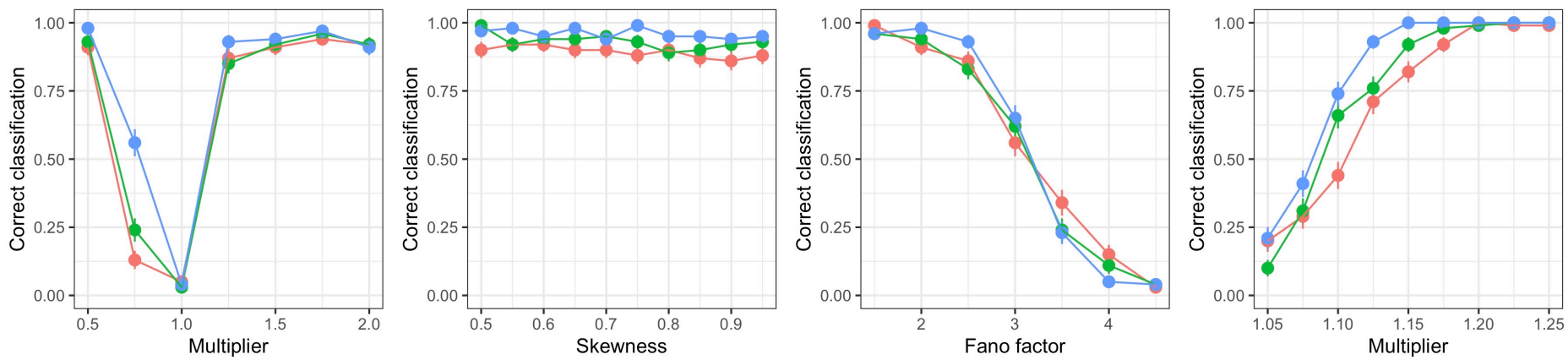
Multiplier

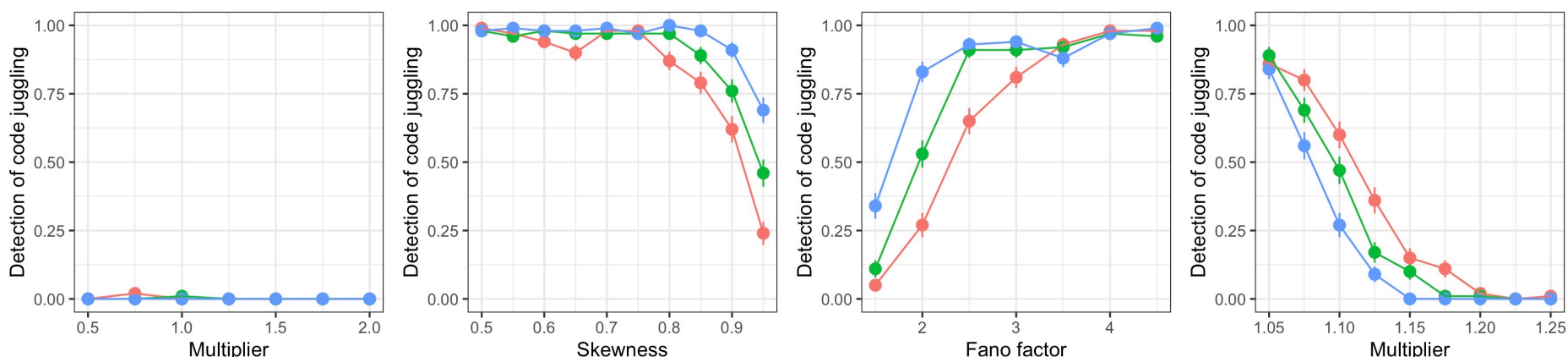
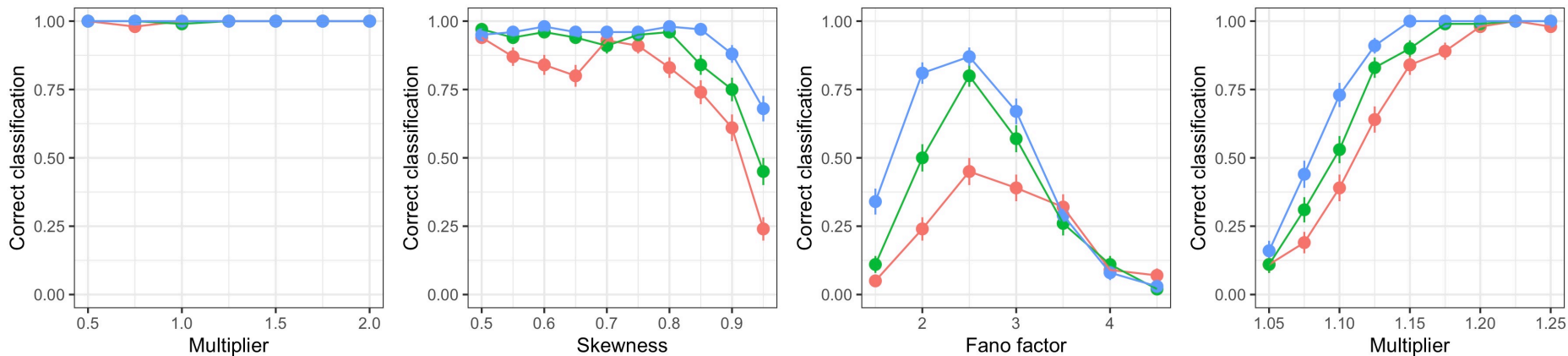
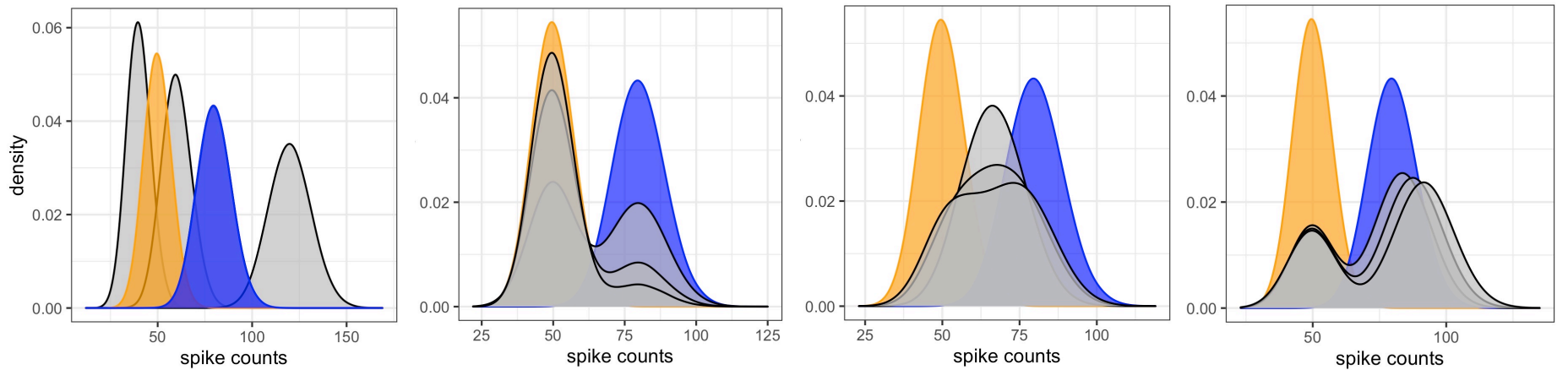




# Performance Assessment

Bounded Uniform



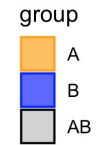
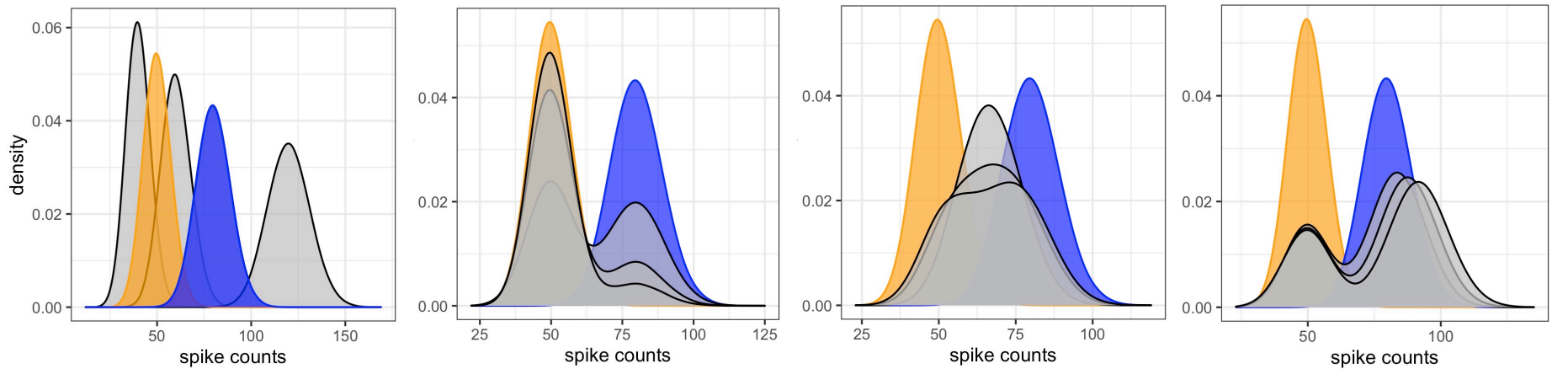


sampsize 20 30 50

# Performance Assessment

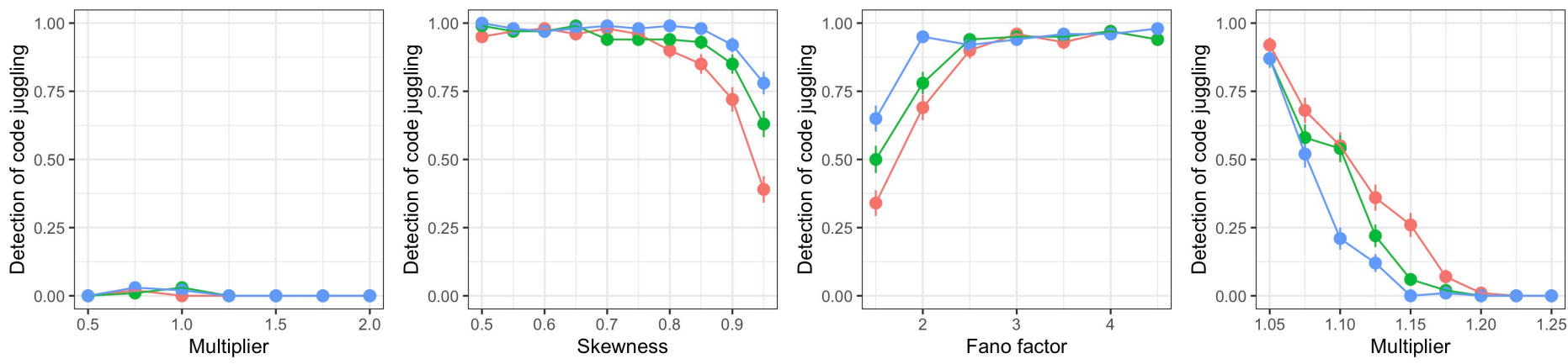
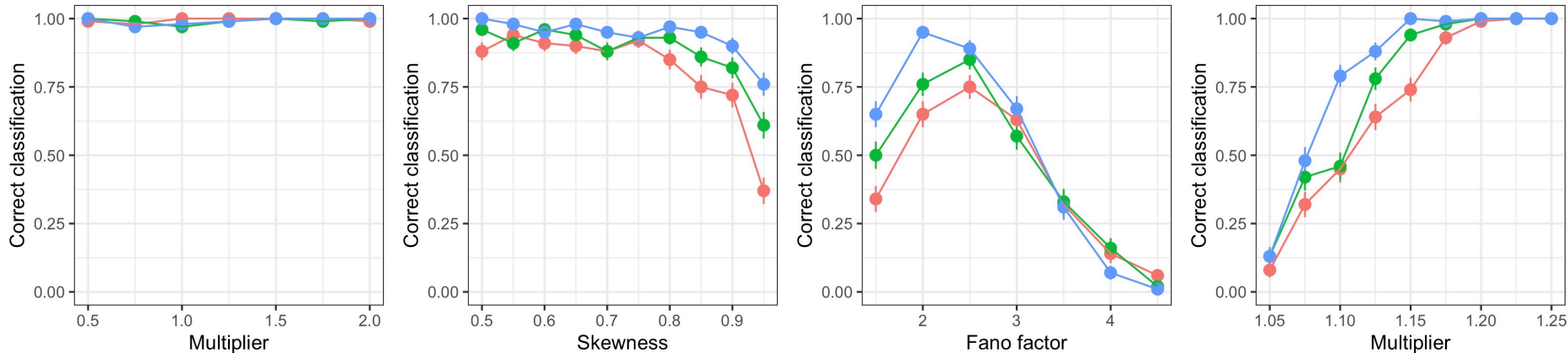
Jeffreys' prior  
(intrinsic Bayes factor)



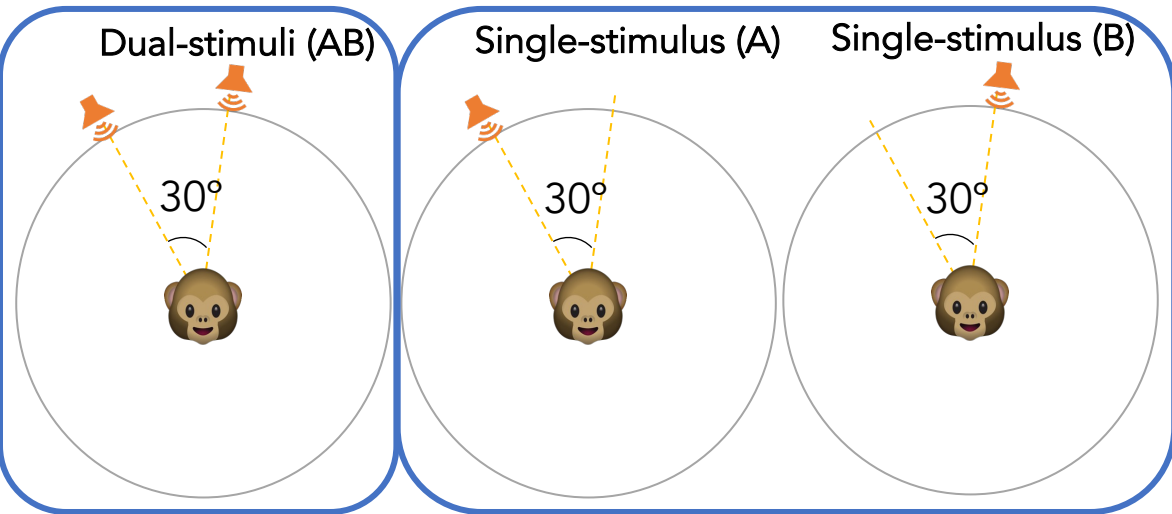


# Performance Assessment

Jeffreys' prior  
(w/ constant 1)



# Application in IC data



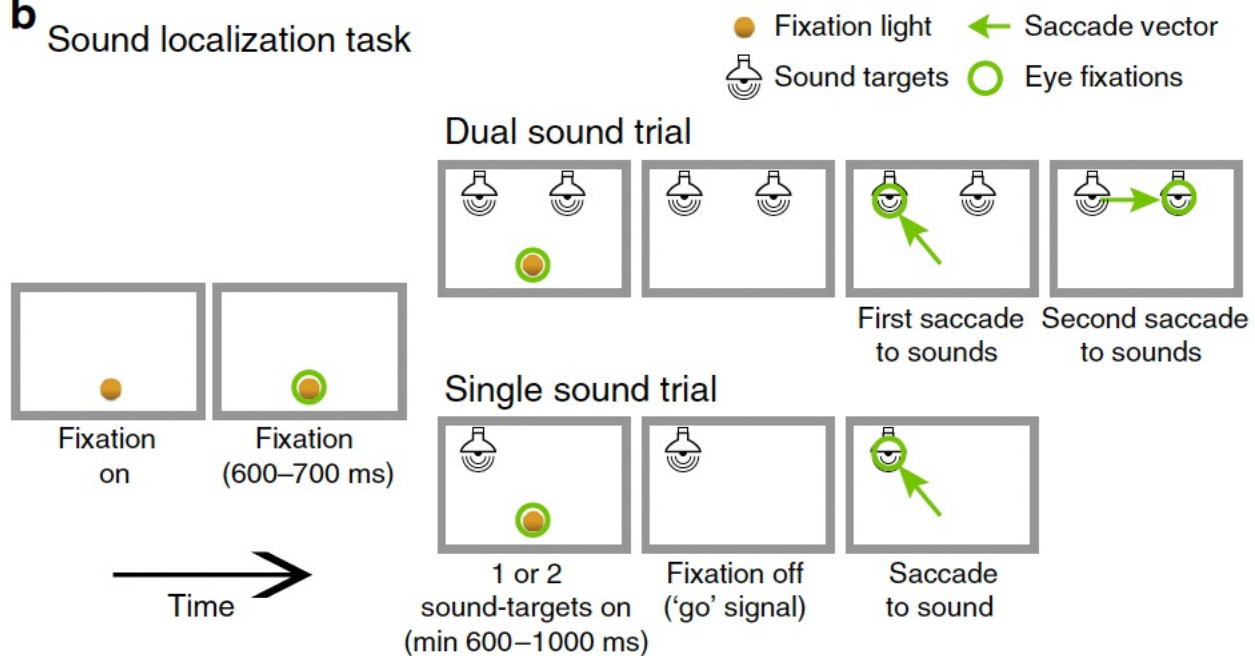
## Experiment design (Caruso et al., 2018)

- Localization task: eye movements to sound (saccades)
- Single cell recording in Inferior colliculus (IC): accurate sound localization

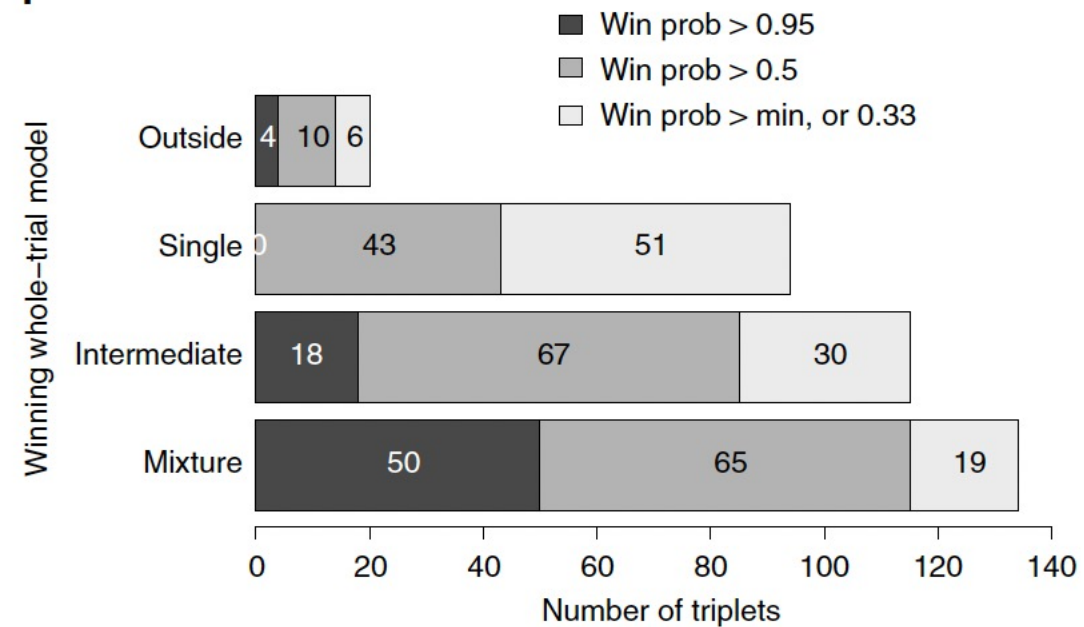
## Results from Caruso et al. (2018)

- Chi-square goodness of fit test: 363 triplets

### b Sound localization task

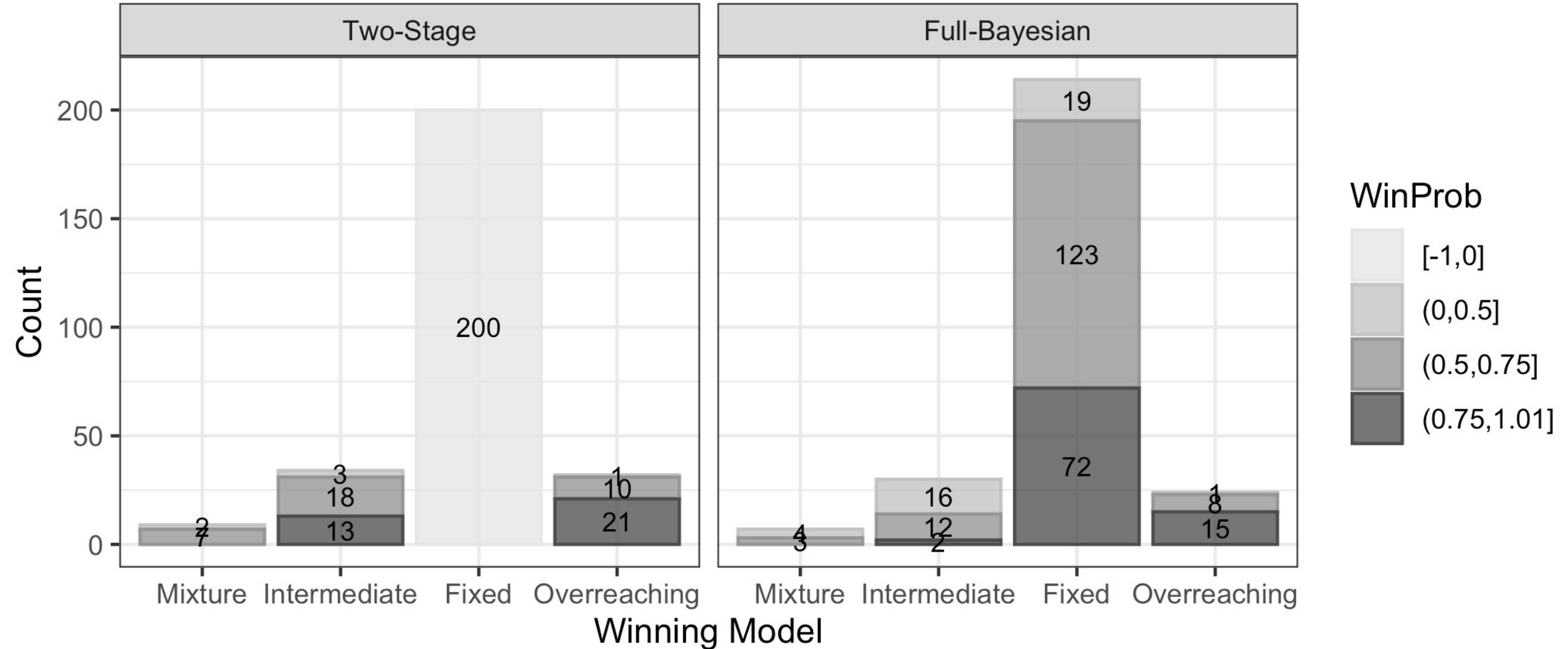


### i

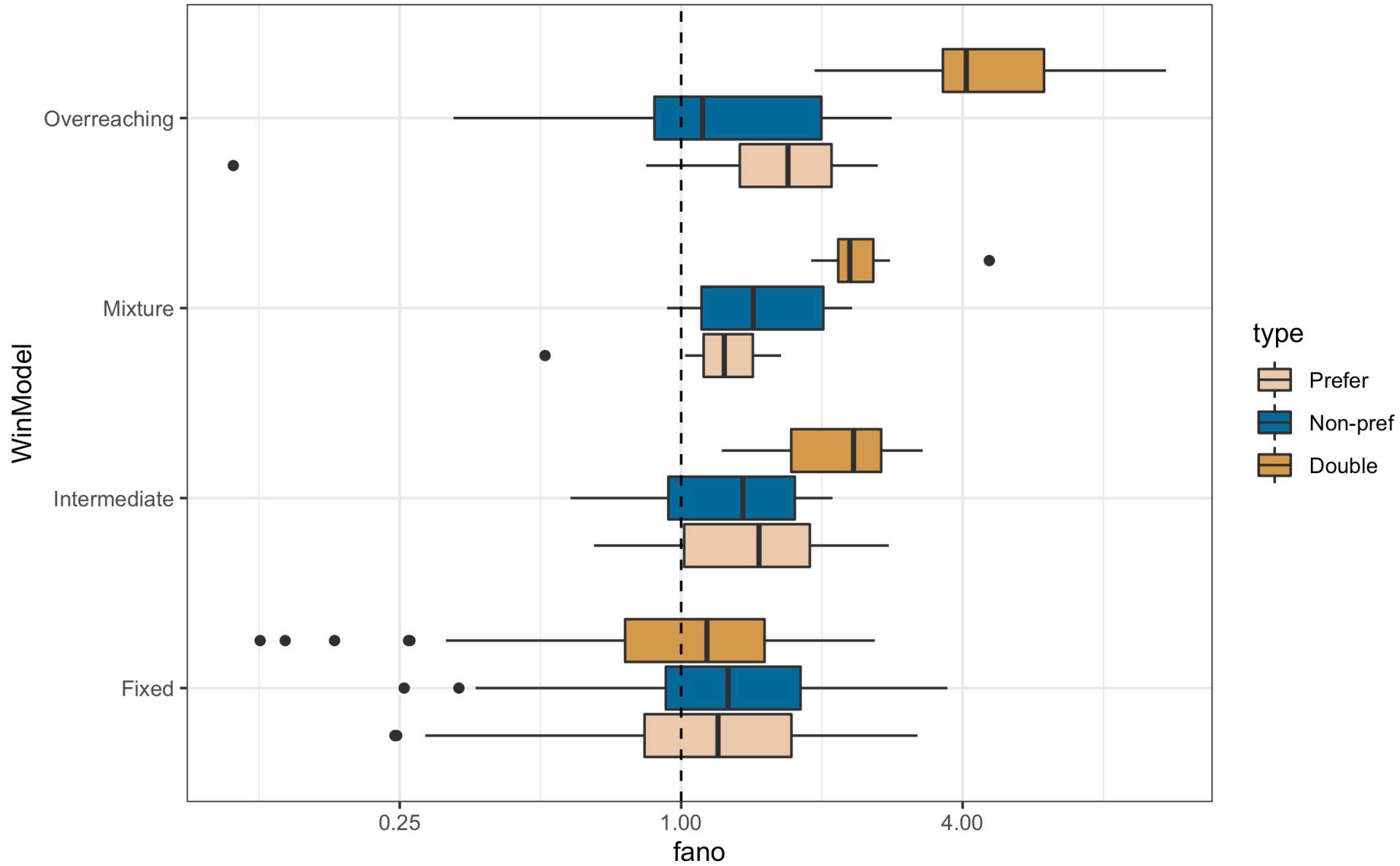


# Results

Preprocessing: Poisson variance test (Brown and Zhao, 2002) ensure Poisson-like distribution for single stimulus

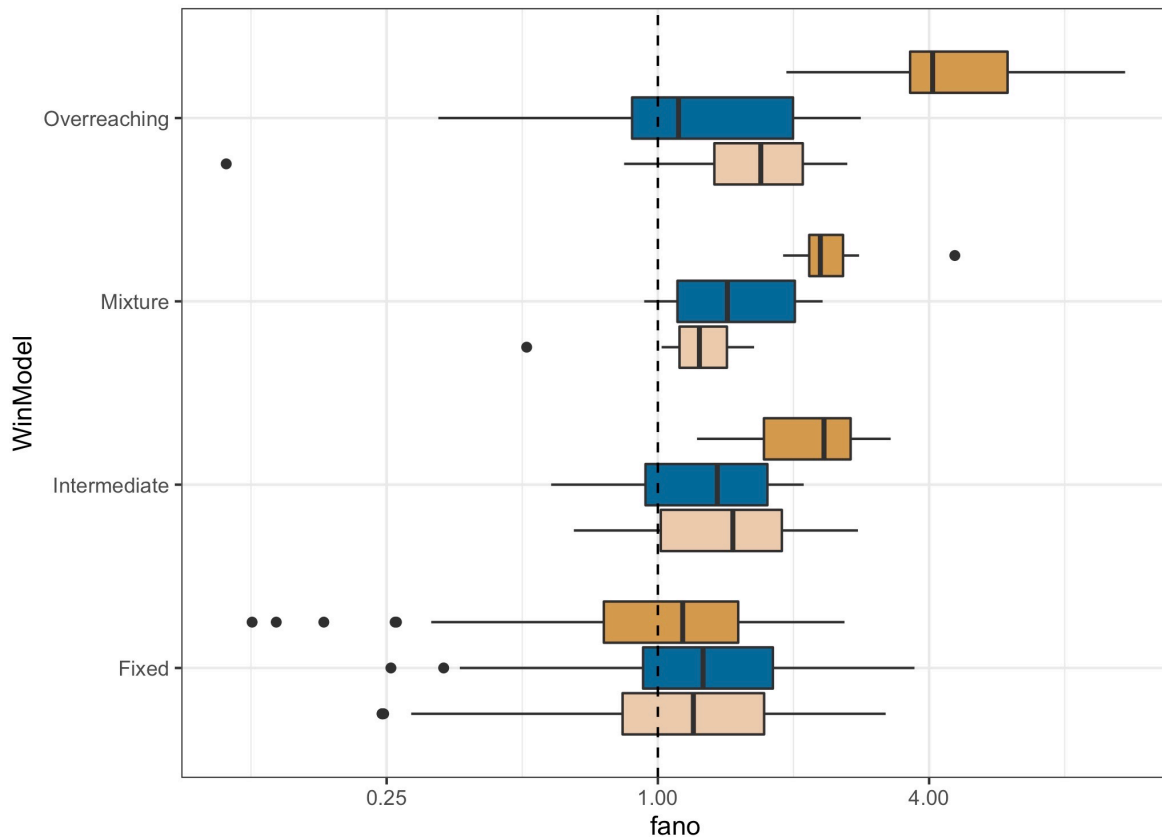


# Results

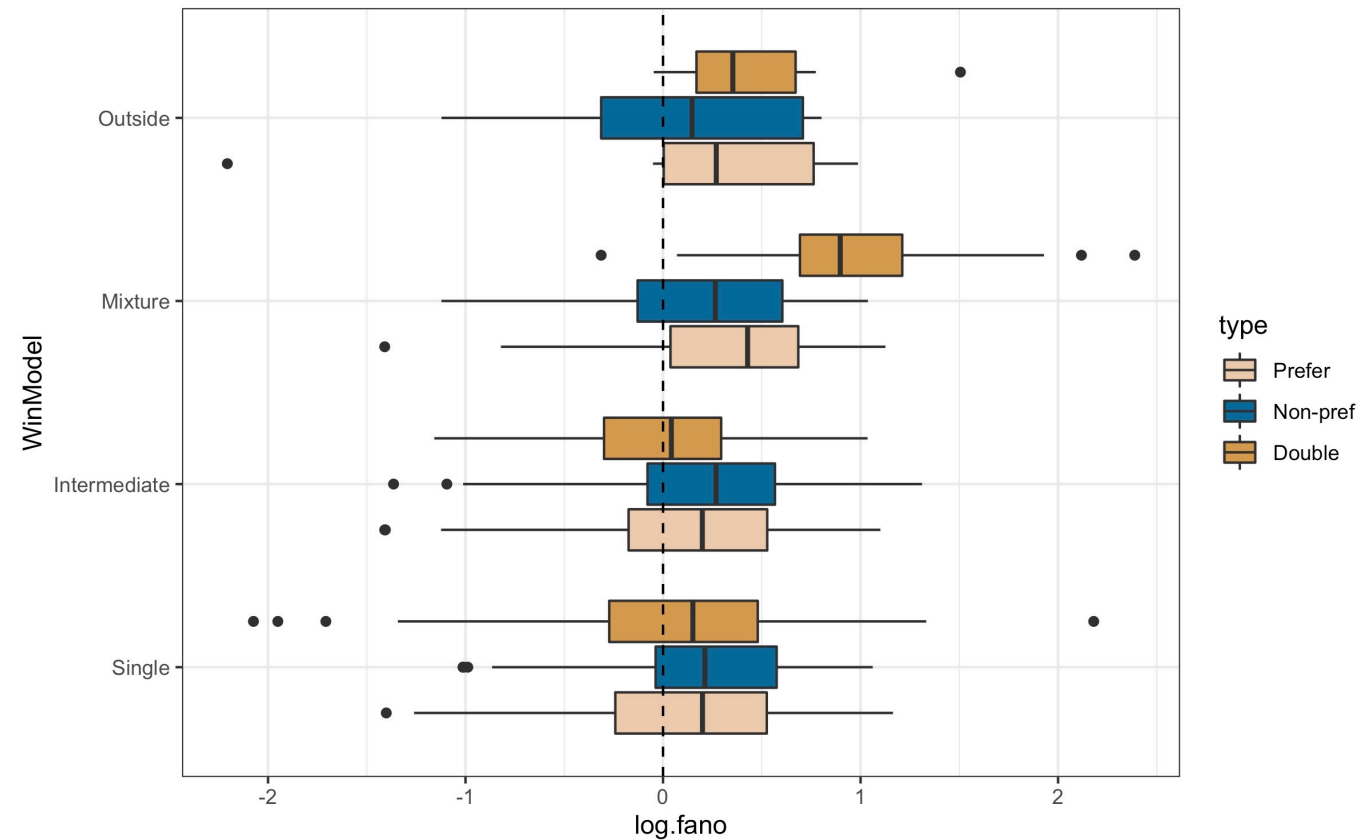


# Results

## New Method



## Old Poisson





**Thank you**

# Reference

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# PRML Algorithm

**Input:** i.i.d observations  $Y_1, \dots, Y_n$

**Output:** marginal likelihood  $L_n(\theta) = \prod_{i=1}^n m_{i-1,\theta}(Y_i)$ , gradient

$\nabla \log L_n(\theta) = \sum_{i=1}^n \nabla \log m_{i-1,\theta}(Y_i)$  and mixing density  $f_{n,\theta}$

**Initialize:**  $f_{0,\theta} \in \mathbb{F}$  of  $f$  (usually uniform); compute  $\nabla f_{0,\theta}(u)$ ;  
weights  $w_1, \dots, w_n \in (0, 1)$  (usually  $w_i = (i + 1)^{-\gamma}$ )

**for**  $i = 1, \dots, n$  **do**

$$m_{i-1,\theta}(Y_i) = \int p(Y_i|\theta, u') f_{i-1,\theta}(u') d\mu(u') \quad (13)$$

$$f_{i,\theta}(u) = (1 - w_i) f_{i-1,\theta}(u) + w_i \frac{p(Y_i|\theta, u) f_{i-1,\theta}(u)}{m_{i-1,\theta}(Y_i)} \quad (14)$$

For  $\nabla f_{i,\theta}(u)$ :

$$G(\theta, u) = p(Y_i|\theta, u) \nabla f_{i-1,\theta}(u) + \nabla p(Y_i|\theta, u) f_{i-1,\theta}(u) \quad (15)$$

$$\nabla \log m_{i-1,\theta}(Y_i) = \frac{\int G(\theta, u) d\mu(u)}{m_{i-1,\theta}(Y_i)} \quad (16)$$

$$\nabla f_{i,\theta}(u) = (1 - w_i) \nabla f_{i-1,\theta}(u) + w_i \left\{ \frac{G(\theta, u) - p(Y_i|\theta, u) f_{i-1,\theta}(u) \nabla \log m_{i-1,\theta}(Y_i)}{m_{i-1,\theta}(Y_i)} \right\} \quad (17)$$

**end**